

RESEARCH

Open Access



Comparing risk-adjusted inpatient fall rates internationally: validation of a risk-adjustment model using multicentre cross-sectional data from hospitals in Switzerland and Austria

Niklaus S. Bernet^{1*}, Irma H. J. Everink², Sabine Hahn¹, Silvia Bauer³ and Jos M. G. A. Schols²

Abstract

Background Inpatient falls in hospitals are an acknowledged indicator of quality of care. International comparisons could highlight quality improvement potential and enable cross-national learning. Key to fair cross-national comparison is the availability of a risk adjustment model validated in an international context. This study aimed to 1) ascertain that the variables of the inpatient fall risk adjustment model do not interact with country and thus can be used for risk adjustment, 2) compare the risk of falling in hospitals between Switzerland and Austria after risk adjustment.

Methods The data on inpatient falls from Swiss and Austrian acute care hospitals were collected on a single measurement day in 2017, 2018 and 2019 as part of an international multicentre cross-sectional study. Multilevel logistic regression models were used to screen for interaction effects between the patient-related fall risk factors and the countries. The risks of falling in hospital in Switzerland and in Austria were compared after applying the risk-adjustment model.

Results Data from 176 hospitals and 43,984 patients revealed an inpatient fall rate of 3.4% in Switzerland and 3.9% in Austria. Two of 15 patient-related fall risk variables showed an interaction effect with country: Patients who had fallen in the last 12 months ($OR\ 1.49, 95\%\ CI\ 1.10-2.01, p=0.009$) or had taken sedatives/psychotropic medication ($OR\ 1.40, 95\%\ CI\ 1.05-1.87, p=0.022$) had higher odds of falling in Austrian hospitals. Significantly higher odds of falling were observed in Austrian ($OR\ 1.38, 95\%\ CI\ 1.13-1.68, p=0.002$) compared to Swiss hospitals after applying the risk-adjustment model.

Conclusions Almost all patient-related fall risk factors in the model are suitable for a risk-adjusted cross-country comparison, as they do not interact with the countries. Further model validation with additional countries is warranted, particularly to assess the interaction of risk factors “fall in the last 12 months” and “sedatives/psychotropic medication intake” with country variable. The study underscores the crucial role of an appropriate risk-adjustment model in ensuring fair international comparisons of inpatient falls, as the risk-adjusted, as opposed to the non-risk-adjusted country comparison, indicated significantly higher odds of falling in Austrian compared to Swiss hospitals.

Keywords Accidental falls, Benchmarking, Cross-sectional studies, Hospitals, Inpatients, Quality assurance, Quality of health care, Multilevel analysis, Risk adjustment

*Correspondence:

Niklaus S. Bernet
niklausstefan.bernet@bfh.ch

Full list of author information is available at the end of the article



© The Author(s) 2024. **Open Access** This article is licensed under a Creative Commons Attribution 4.0 International License, which permits use, sharing, adaptation, distribution and reproduction in any medium or format, as long as you give appropriate credit to the original author(s) and the source, provide a link to the Creative Commons licence, and indicate if changes were made. The images or other third party material in this article are included in the article's Creative Commons licence, unless indicated otherwise in a credit line to the material. If material is not included in the article's Creative Commons licence and your intended use is not permitted by statutory regulation or exceeds the permitted use, you will need to obtain permission directly from the copyright holder. To view a copy of this licence, visit <http://creativecommons.org/licenses/by/4.0/>. The Creative Commons Public Domain Dedication waiver (<http://creativecommons.org/publicdomain/zero/1.0/>) applies to the data made available in this article, unless otherwise stated in a credit line to the data.

Introduction

International comparisons of quality of care outcomes have gained popularity in recent years [1, 2] and have proven to be an effective strategy for cross-national learning and the promotion of quality development [3]. By making reference values from other countries transparent, external benchmarking can highlight areas where health services are not performing as well as they could and therefore have potential for improvement [3–5].

A notable initiative in this realm, for example, is the Health Care Quality Indicators (HCQI) project led by the Organisation for Economic Co-operation and Development (OECD) [6–8]. This project was launched in 2003 to develop and publish indicators for international comparisons of health care quality, enabling benchmarking and mutual learning [4, 6, 7]. The indicators cover aspects such as acute care (e.g., 30-day mortality after acute myocardial infarction) and patient safety (e.g., post-operative complications) [6] and are published on OECD Health Statistics [9] and in the biennial OECD series “Health at a Glance” [10].

Despite the recognised potential and evidence that there are differences in quality of care outcomes within countries, but also between countries [4, 11], international comparisons are still rarely made [1]. This is predominantly attributable to the following four methodological challenges. First, despite recent developments such as defined core outcome sets [12], an internationally standardised data collection procedure is often lacking. This results in the absence of uniformly collected data [3, 13]; for example, if the information is based on primary data collection in one country or on routine data in another country [14, 15]. Second, internationally standardised definitions are frequently not applied [3, 16]. This can lead to systematic differences between countries, e.g., if a fall event is defined and calculated differently in the countries concerned, and the numerators are therefore treated differently [15]. Third, data protection regulations in different countries restrict or prevent the exchange of data across national borders and thus the use of data in an international context [17], especially when data with patient identifiers are required [1]. Fourth, in addition to data availability and data quality, the aspect of risk adjustment (or case-mix adjustment) is another challenge in the effort to enable fair and meaningful comparisons of healthcare quality outcomes (inter)nationally [1].

Differences in outcome measures between different groups, e.g., health care providers or countries, can be explained by various factors, such as differences in quality of care, in definitions, in data quality, in case mix and by chance [15]. Risk adjustment is a method that aims to control for differences in case mix. It is defined as “the process of statistically accounting for differences in

patient mix that affect health care outcomes” [18]. When applying risk adjustment methods, i.e., adjusting for factors outside the direct control of health care providers, there is more certainty that differences in health care outcomes can be attributed to differences in the quality of care provided [11]. Thus, the process of risk adjustment is often paraphrased as “levelling the playing field” for a fair comparison [3, 19].

A key requirement of a risk adjustment model at an international level is that patient-related risk factors in the model have a universal character, i.e., the effects of the risk variables on the outcome should be more or less comparable in all countries studied [1, 20]. In other words, age as a patient-related risk factor should, e.g., be associated with worse outcome in country X to the same extent as in country Y. However, this required so called constant risk relationship is often not present in international comparisons, or at least unclear because it is not investigated or reported [1, 20]. By testing for interaction effects, the presence of constant risk relationships can be investigated [1, 20–22]. If no relevant interaction effects are found in the model, it indicates constant risk relationships between the patient-related risk factors and the grouping variable country, as the risk factors in each country show the same effect on the outcome.

In sum, to ensure a trustworthy international comparison of quality indicators, uniformly standardised and collected data as well as a risk adjustment model validated in an international context with regard to possible interactions must be in place. Meeting these requirements is a considerable challenge. Accordingly, few internationally comparable data and, in particular, ready-to-use risk adjustment models are available. This is especially true if, in contrast to mortality and hospitalisation rates, the quality indicators are predominantly based on primary clinical data, as in the case of the inpatient fall rate.

Inpatient fall rates in the acute care hospital setting are an internationally recognised indicator of quality of care [23, 24]. Inpatient falls are among the most frequently reported adverse events and lead to high healthcare costs, e.g., costs related to injuries and longer hospital stays [25–27]. Despite the consequences of such falls in hospitals, we are not aware of any risk-adjusted international comparisons based on clinical data. Nevertheless, an international comparison of inpatient fall rates would provide an important external reference benchmark for decision makers such as governments, health policy makers and health service managers, as well as for health professionals themselves, to assess possible national quality improvement potentials. Especially since a recently published study by Bernet et al. [28] based on data from the International Prevalence Measurement of Quality of Care (LPZ measurement) [29] showed that the performance of

Swiss acute care hospitals regarding inpatient falls was very homogeneous in a national risk-adjusted comparison. The authors conclude that without a risk-adjusted external reference benchmark, there is no indication of whether lower fall rates could potentially be achieved by national quality improvements or whether high quality is already being provided compared to other countries [28]. Therefore, a risk-adjusted international comparison of inpatient fall rates can provide decision-makers at all levels with the evidence required for resource allocation. The question that needs to be answered is whether it is worth investing additional resources in inpatient fall prevention measures or whether these resources could lead to more far-reaching improvements in another area with greater quality improvement potential. To ensure that an international comparison is fair and can be used as a basis for decision-making, the inpatient fall risk adjustment model must first be validated in an international context, i.e., it must be verified whether there is a constant risk relationship between the patient-related risk factors and the countries in relation to the outcome.

As a recently developed inpatient fall risk adjustment model [28] and uniformly collected data on inpatient falls in acute care hospitals in Switzerland and Austria based on the LPZ measurement are available, we aimed to complement the Swiss data used by Bernet et al. [28] with the corresponding data from Austria in order to validate the inpatient fall risk adjustment model in an international context. Therefore, the aims of this study were:

- 1) to investigate whether the patient-related fall risk variables included in the inpatient fall risk adjustment model by Bernet et al. [28] show a constant risk relationship with outcome by not interacting with country, thus allowing the model to be used for a risk-adjusted comparison of the odds of falling in hospital in Switzerland and Austria.
- 2) to investigate whether the odds of falling in hospital in Switzerland and Austria differ after risk adjustment.

Methods

Study design

In this study, a data analysis was conducted based on data retrieved in the International Prevalence Measurement of Quality of Care (Landelijke Prevalentiemeting Zorgkwaliteit, LPZ measurement [29]). The LPZ measurement is an annual international multicentre cross-sectional study that captures background characteristics of patients, prevalence and prevention measures related to different quality of care indicators such as falls, pressure injuries and malnutrition in various care settings and countries. A more detailed description

on the aim and procedures of the LPZ measurement can be found in Van Nie et al. [29].

This study used data on inpatient falls in Swiss and Austrian acute care hospitals (university hospitals, general hospitals and specialised clinics). The data were collected annually by the hospitals on the second Tuesday in November (14 November 2017, 13 November 2018 and 12 November 2019) using the same standardized LPZ measurement method. Other countries participating in the LPZ measurement could not be included because either no or only very few acute care hospitals participated in their measurements. The present study was conducted and reported in accordance with the RECORD (REporting of studies Conducted using Observational Routinely-collected health Data) Statement [30, 31]. The corresponding checklist used can be found in Additional file 1.

Setting and sample

In Switzerland, all acute care hospitals that have signed up to the national quality contract (approximately 97% of Swiss acute care hospitals) are obliged to participate in the LPZ measurement. These hospitals were annually contacted by the national coordinator of Switzerland (Bern University of Applied Sciences) by e-mail and asked to register for the measurement. The national coordinator for the LPZ measurement in Austria (Medical University of Graz) annually invites all acute care hospitals by email to participate in the measurement on a voluntary basis. Participation in the LPZ measurements was subject to a fee for the hospitals. With minor differences between the countries and slight adaptations over the years, the amount the hospitals paid was calculated annually based on a fixed amount per hospital of around 500 euros and a variable amount of around 3 euros per bed. In Switzerland, however, the costs for the hospitals were covered by the Swiss Association for Quality Development in Hospitals and Clinics (ANQ).

On the measurement day of the LPZ measurement, all inpatients who were 18 years or older and from whom a verbal (Switzerland) or written (Austria) informed consent was available were included. Patients were excluded from the LPZ measurement if they were not physically present on the ward during the measurement (e.g., due to therapies or surgeries) or if informed consent was not available from the patient (e.g., comatose or terminal condition, cognitive impairment, language barrier) or their legal representative.

All patients from psychiatric or closed geriatric wards were excluded from this analysis, as these ward types participated in Austria but not in Switzerland.

Data collection

Data collection in the LPZ measurement follows a standardised procedure, which is described in detail in the internationally defined measurement manual and is available to all people involved in the measurement in the respective language(s) of the participating country. Following the train-the-trainer principle, the national coordinators of Switzerland and Austria organised an annual training event for the measurement coordinators of each participating hospital prior to the measurement. During these trainings the measurement procedure including the inclusion and exclusion criteria, the definitions, the questionnaire and the web-based data entry programme were explained in detail. The training materials used were also made available to the hospital coordinators so that they could use them to train their clinical measurement teams.

The clinical measurement teams each consisted of two nurses (registered nurses or at least nurses who have assessment responsibilities in everyday clinical practice, one of whom works on the ward being surveyed and the other one on a different ward). During the measurement, the clinical measurement teams collected the data directly from the patient or from medical records, where permissible, according to the measurement manual and questionnaire (e.g., patient characteristics) [29].

Responses were either recorded on paper questionnaires and then transferred to the data entry programme or, alternatively, were entered directly into the data entry programme on the LPZ website. The data entry programme is designed to guide participants intuitively through the questionnaire, to display warnings for unanticipated data entries, and to prevent missing values by not allowing a questionnaire to be completed until all mandatory data have been entered. Furthermore, the transfer of the data and the data quality were determined by the operator of the LPZ data entry programme using a data profile. This involves a systematic search for error patterns in the datasets (data that should not be present; conspicuous or missing values). In addition, the national data sets are checked by the national coordinators for implausible cases, missing data and conspicuous values (e.g., very high age entries). Potentially incorrect data entries are verified together with the hospitals concerned and, where indicated, corrected by the hospitals directly in the data entry programme [32].

Measures

The data were collected using the LPZ 2.0 instrument, which includes validated measures to assess the prevalence of various quality of care indicators, and is continuously evaluated and revised as needed by experts from an international research group [29]. The LPZ 2.0

instrument measures general patient background characteristics such as age, gender, diagnoses, etc., and quality of care indicator specific information at institution, ward and patient level. Available information that was not relevant for this study, such as information on other quality indicators (e.g., pressure injuries, malnutrition), was not included in the analysis.

The outcome variable of this study is whether or not a patient fell in hospital. Following the definition of the Kellogg International Work Group on the Prevention of Falls by the Elderly [33], a fall is defined in the LPZ measurement as any unintentional change in position that results in the client coming to rest on the ground or other lower level, regardless of the reason [28, 29]. Inpatient falls were retrospectively assessed with the question: “Has the client fallen in the last 30 days in this institution?” (yes/no). All other variables considered in this study are patient-related fall risk factors that were identified as relevant based on the inpatient fall risk adjustment model developed by Bernet et al. [28]. Table 1 gives an overview of the variables used in the present study and in the inpatient fall risk adjustment model of Bernet et al. [28].

Data analysis

The data analysis was divided into seven main steps, which are described in detail below (see also Additional file 2 for a summarised overview of the main analysis steps).

First, the complete Swiss dataset used for the development of the inpatient fall risk adjustment model in the study by Bernet et al. [28] was supplemented with the complete Austrian datasets of the 2017, 2018 and 2019 measurements.

Second, to obtain comparable data sets between Switzerland and Austria and to promote accurate estimates in the context of a multilevel model, all hospitals with fewer than 50 participants over the three measurement years were excluded from the analysis, analogous to the procedure described in Bernet et al. [28].

Thirdly, for sample description, the number of hospitals and the patient-related fall risk factors according to Table 1 were described with frequencies and percentages. In addition, Pearson's chi-square test was used to examine whether patient-related fall risk factors differed significantly between the two countries. As a baseline model to illustrate a non-risk-adjusted comparison of inpatient fall rates across Switzerland and Austria, a two-level random intercept logistic regression model was calculated, with hospitals modelled as a random effect and country as a fixed effect. The corresponding odds ratio (OR) and 95% confidence interval (95% CI) were reported.

Fourth, to answer the question of whether there are interaction effects between the patient-related fall risk

Table 1 Variables included in the inpatient fall risk adjustment model and therefore in the present study

Outcome variable	Answer options
Has the client fallen in the last 30 days in this institution	no [0] yes [1]
Characteristics/patient-related fall risk factors	Answer options
Age (in years) / Age groups	Scale / 18–64 years [0] 65–74 years [1] 75–84 years [2] ≥ 85 years [3]
Sex	male [0] female [1]
Care Dependency Scale (CDS)	care independent (70–75) [0] to a great extent independent (60–69) [1] partially dependent (45–59) [2] to a great extent dependent (25–44) [3] completely depend- ent (15–24) [4]
Fall in the last 12 months before hospital admission	no [0] yes [1]
Sedative/psychotropic medication intake	no [0] yes [1]
Surgical procedure within 14 days prior to measurement	no [0] yes [1]
ICD-10—Mental and behavioural disorders [yes]	no [0] yes [1]
ICD-10—Neoplasms	no [0] yes [1]
ICD-10—Diseases of the blood and blood-forming organs	no [0] yes [1]
ICD-10—Certain infectious and parasitic diseases	no [0] yes [1]
ICD-10—Factors influencing health status	no [0] yes [1]
ICD-10—Diseases of the nervous system	no [0] yes [1]
ICD-10—Endocrine, nutritional and metabolic diseases	no [0] yes [1]
ICD-10—Diseases of the musculoskeletal system	no [0] yes [1]
ICD-10—Diseases of the ear [yes]	no [0] yes [1]

The Care Dependency Scale consists of 15 items (e.g., food and drink, continence, mobility). The sum score ranges from 15 to 75 points [34, 35] and can be categorized into five categories: Completely dependent on care from others, to a great extent dependent, partially dependent, to a great extent independent and care independent [36, 37]

ICD-10 International statistical classification of diseases and related health problems 10th revision [38]

factors of the inpatient fall risk adjustment model of Bernet et al. [28] and the country, the inpatient fall risk adjustment model was applied to the merged Swiss and Austrian data within the framework of a two-level random intercept logistic regression model. In this, the hospitals are modelled as random effect and the

patient-related fall risk factors according to Table 1 as fixed effects. In addition, the interaction terms between the patient-related fall risk factors and the country are considered as fixed effects in the model (hereafter denoted as “Model 1”). Following Moger and Pelto [1], Mohammed et al. [20], Goodacre et al. [21] the

interaction terms were used to screen whether the proposed inpatient fall risk adjustment model fulfils the basic requirement of including universal fall risk factors, i.e., whether they are constantly associated with a higher or lower inpatient fall risk regardless of the country and can therefore safely be used for a risk-adjusted international comparison of inpatient falls in Switzerland and Austria. If interaction effects are present, they may indicate a non-constant risk relationship and thus the danger of the so-called constant risk-fallacy [22]. The constant risk fallacy means that, contrary to the intention of risk adjustment to reduce bias in comparisons, bias may instead increase [1, 20, 22]. Significant interaction effects, however, do not provide any information as to the causes of the non-constant risk relationships found, and consequently whether these are problematic for a risk-adjusted country comparison (or whether the non-constant risk relationships are instead due to differences in the quality of care provided). If there are significant interaction effects, however, this is an indication that further attention is needed. The inpatient fall risk adjustment model did not converge initially, which is not uncommon in complex modelling situations but may lead to untrustworthy model estimates [39]. As the convergence issue was related to the continuous variable “age” in the model, we decided to transform the variable into a categorical variable in order to reduce model complexity. Following Heikkilä et al. [40] and Thoman et al. [41] the subsequent four age categories were created: 18–64 years, 65–74 years, 75–84 years and older than 84 years. By introducing “age” as a categorical variable in the inpatient fall risk adjustment model, the convergence problems could be solved. To test for a possible measurement year effect, we recalculated “Model 1” by including the measurement year as a control variable. The measurement year was not significant in the model and the Akaike Information Criterion (AIC) value was higher than in the initial “Model 1”. Therefore, only the initial model (“Model 1”) without the control variable measurement year was subsequently reported.

Fifth, in the context of a model development process, a reduced two-level random intercept logistic regression model (hereafter denoted as “Model 2”) was calculated to capture the main interaction effects. In “Model 2”, in addition to the patient-related fall risk factors, only those 2×2 interaction terms were included in the model that had a *p*-value smaller than 0.16 in “Model 1”. This less stringent threshold (compared to a *p*-value < 0.05) was chosen so that important interaction effects would not be overlooked due to lack of power or stochastic variability and because this value roughly corresponds to a variable selection based on the AIC [42]. In addition, to decide whether the interaction effects of the variables age and CDS, which have several factor levels, are

relevant overall, AIC was used to compare “Model 1” with a model without the interaction term Age*Country or CDS*Country. If the AIC value in the model without the corresponding interaction term was lower than in “Model 1”, the corresponding interaction term was not taken into account in “Model 2”.

Sixth, to clarify whether the inpatient fall risk adjustment model accounts equally for key patient-related fall risk variables in Switzerland and Austria, we applied the model separately to the Swiss and Austrian datasets. We subsequently compared the performance of the model in terms of discrimination using receiver operating characteristic (ROC) curve analysis.

Seventh, to answer the question of whether the risk of falling in hospital in Switzerland and Austria differ after risk adjustment, we applied the inpatient fall risk adjustment model of Bernet et al. [28] to the merged Swiss and Austrian data within the framework of a two-level random intercept logistic regression model. The hospitals are modelled as random effect and the patient-related fall risk factors according to Table 1 as fixed effects.

Data preparation and descriptive analyses were carried out with IBM® SPSS® Statistics (version 28). The multi-level logistic regression models and the ROC curves were built and analysed with the statistics programme R, version 4.1.0 [43], and the packages *lme4* [44], *sjplot* [45], *effects* [46, 47] and *pROC* [48]. The statistical significance level was set at *p*-value below 0.05 in all analyses.

Results

A total of 176 hospitals with 43,984 patients who participated in the 2017, 2018 and 2019 measurements in Switzerland and Austria were included in the analysis. Of these, 138 (78.4%) hospitals were from Switzerland and 38 (21.6%) from Austria. Thus, the majority of the participants were hospitalised in Swiss hospitals (81.8%, *n*=35,998). The distribution of participants across hospital types was similar in the two countries, with the largest proportion of participants coming from general hospitals (Switzerland 73.9%, *n*=26,590; Austria 70.3%, *n*=5,611), followed by university hospitals (Switzerland 19.4%, *n*=6,982; Austria 26.1%, *n*=2,088) and specialised hospitals (Switzerland 6.7%, *n*=2,426; Austria 3.6%, *n*=287). Almost half of the participants were female (49.5%, *n*=21,787), aged between 18–74 years (60.9%, *n*=26,770) and care independent according to the CDS (55.2%, *n*=24,273). Nearly one third of participants had experienced a fall in the last 12 months prior to hospital admission (29.7%, *n*=13,073). A similar proportion had been exposed to current (in the past 24 h), periodic (irregular) and/or continuous (regular) intake of sedatives or psychotropic medication (35.0%, *n*=15,377). Further, 40.8% (*n*=17,954) of all participants had undergone a surgical

Table 2 Descriptive and bivariate comparison of patient-related fall risk factors between Switzerland and Austria

Patient-related fall risk factors	Switzerland (n = 35,998)		Austria (n = 7,986)		p-value
	n	%	n	%	
Sex [female]	17,669	49.1	4,118	51.6	< 0.001
Age groups					
18–64 years	13,738	38.2	2,981	37.3	< 0.001
65–74 years	8,186	22.7	1,865	23.4	
75–84 years	8,913	24.8	2,133	26.7	
≥ 85 years	5,161	14.3	1,007	12.6	
Care dependency (CDS)					
care independent [70–75]	19,247	53.5	5,026	62.9	< 0.001
to a great extent independent [60–69]	8,807	24.5	1,320	16.5	
partially dependent [45–59]	5,218	14.5	849	10.6	
to a great extent dependent [25–44]	2,083	5.8	516	6.5	
completely dependent [15–24]	643	1.8	275	3.4	
Fall in the last 12 months [yes]	11,131	30.9	1,942	24.3	< 0.001
Sedative/psychotropic medications intake [yes]	12,928	35.9	2,449	30.7	< 0.001
Surgical procedure within 14 days prior to measurement [yes]	15,885	44.1	2,069	25.9	< 0.001
ICD-10—Mental and behavioural disorders [yes]	6,932	19.3	1,093	13.7	< 0.001
ICD-10—Neoplasms [yes]	7,895	21.9	1,557	19.5	< 0.001
ICD-10—Diseases of the blood and blood-forming organs [yes]	6,074	16.9	709	8.9	< 0.001
ICD-10—Certain infectious and parasitic diseases [yes]	4,788	13.3	585	7.3	< 0.001
ICD-10—Factors influencing health status [yes]	2,513	7.0	691	8.7	< 0.001
ICD-10—Diseases of the nervous system [yes]	5,172	14.4	1,072	13.4	0.029
ICD-10—Endocrine, nutritional and metabolic diseases [yes]	12,617	35.0	2,733	34.2	0.161
ICD-10—Diseases of the musculoskeletal system [yes]	14,626	40.6	2,197	27.5	< 0.001
ICD-10—Diseases of the ear [yes]	959	2.7	145	1.8	< 0.001

CDS Care dependency scale, ICD-10 International statistical classification of diseases and related health problems 10th revision

Significant differences between countries according to the Pearson's chi-square test results are highlighted in bold

procedure in the 14 days prior to the measurement. The participants were most frequently affected by a diagnosis from the ICD-10 group “Diseases of the musculoskeletal system” (38.2%, $n = 16,823$), followed by “Endocrine, nutritional and metabolic diseases” (34.9%, $n = 15,350$) and “Neoplasms” (21.5%, $n = 9,452$).

When comparing the patient-related fall risk factors between the two countries (see Table 2 and Additional file 3 for a graphic overview), a significant difference was found in almost all variables. The most prominent differences were the higher proportion of care independent participants in Austria, as well as a lower proportion of participants in Austria who had had a fall in the last 12 months. Also significantly fewer participants in Austria were taking sedatives and/or psychotropic medication and had had a surgical procedure in the 14 days prior to the measurement. Of the nine ICD-10 diagnosis groups reported, the percentages of Austrian participants with these diagnoses were significantly lower than in Switzerland for seven, and significantly higher for one of the ICD-10 diagnosis groups.

The outcome analyses showed that a total of 1,551 participants experienced an inpatient fall. This corresponds to an overall inpatient fall rate of 3.5%. The inpatient fall rate by country was 3.4% ($n = 1,239$) in Switzerland and 3.9% ($n = 312$) in Austria. The non-risk-adjusted country comparison (baseline model that, in addition to the random effect, only takes the country variable into account as a fixed effect) showed that the odds of falling in hospital do not differ between Switzerland and Austria (OR 0.99, 95% CI 0.77–1.28, $p = 0.961$).

Interactions between patient-related fall risk factors and country

In “Model 1” (see Table 3 for a description) the interaction effects between the 15 patient-related fall risk factors and the two countries were taken into account. There was a significant interaction effect between the variable “country” and “having a fall in the last 12 months before hospital admission” on the outcome measure inpatient falls (OR 1.33, 95% CI 1.00–1.76, $p = 0.049$). In addition, two interaction effects were found with a p -value of less

Table 3 Description of “Model 1” and “Model 2”¹

Predictors	Model 1			Model 2		
	OR	95% CI	p-value	OR	95% CI	p-value
CDS [care independent (70–75)]	Ref			Ref		
CDS [to a great extent independent (60–69)]	2.20	1.86 – 2.61	<0.001	2.24	1.93 – 2.61	<0.001
CDS [partially dependent (45–59)]	3.08	2.58 – 3.68	<0.001	2.97	2.53 – 3.48	<0.001
CDS [to a great extent dependent (25–44)]	3.55	2.88 – 4.38	<0.001	3.31	2.75 – 3.99	<0.001
CDS [completely dependent (15–24)]	1.95	1.34 – 2.83	0.001	1.96	1.43 – 2.68	<0.001
Fall in the last 12 months [yes]	2.17	1.91 – 2.46	<0.001	2.20	1.95 – 2.50	<0.001
Sedative/psychotropic medication intake [yes]	1.75	1.54 – 1.98	<0.001	1.78	1.57 – 2.02	<0.001
Age [18–64 years]	Ref			Ref		
Age [65–74 years]	1.35	1.12 – 1.62	0.002	1.45	1.23 – 1.71	<0.001
Age [75–84 years]	1.60	1.34 – 1.90	<0.001	1.59	1.36 – 1.86	<0.001
Age [≥ 85 years]	1.55	1.28 – 1.88	<0.001	1.56	1.31 – 1.85	<0.001
ICD-10—Mental and Behavioural disorders [yes]	1.55	1.36 – 1.77	<0.001	1.52	1.35 – 1.71	<0.001
ICD-10—Neoplasms [yes]	1.26	1.10 – 1.44	0.001	1.21	1.07 – 1.37	0.002
ICD-10—Diseases of the blood and blood-forming organs [yes]	1.23	1.07 – 1.41	0.003	1.25	1.10 – 1.42	0.001
ICD-10—Certain infectious and parasitic diseases [yes]	1.19	1.02 – 1.38	0.029	1.15	1.00 – 1.33	0.055
ICD-10—Factors influencing health status [yes]	1.21	0.99 – 1.46	0.059	1.22	1.03 – 1.45	0.023
ICD-10—Diseases of the nervous system [yes]	1.14	0.99 – 1.32	0.068	1.14	1.00 – 1.30	0.045
ICD-10—Endocrine, nutritional and metabolic diseases [yes]	1.13	1.00 – 1.28	0.047	1.12	1.01 – 1.25	0.035
ICD-10—Diseases of the musculoskeletal system [yes]	0.91	0.80 – 1.03	0.116	0.89	0.80 – 1.00	0.050
Surgical procedure within 14 days prior to measurement [yes]	0.82	0.72 – 0.93	0.003	0.81	0.71 – 0.92	0.002
Sex [female]	0.79	0.70 – 0.89	<0.001	0.76	0.68 – 0.85	<0.001
ICD-10—Diseases of the ear [yes]	0.69	0.48 – 0.99	0.043	0.69	0.49 – 0.97	0.031
Country [Austria]	1.36	0.88 – 2.12	0.168	1.19	0.89 – 1.60	0.250
CDS [to a great extent independent (60–69)]:Country [Austria]	1.14	0.79 – 1.64	0.475 ²			
CDS [partially dependent (45–59)]:Country [Austria]	0.82	0.55 – 1.22	0.326 ²			
CDS [to a great extent dependent (25–44)]:Country [Austria]	0.72	0.46 – 1.15	0.172 ²			
CDS [completely dependent (15–24)]:Country [Austria]	1.04	0.52 – 2.06	0.920 ²			
Fall in the last 12 months [yes]:Country [Austria]	1.33	1.00 – 1.76	0.049	1.25	0.96 – 1.64	0.102
Sedative/psychotropic medication intake [yes]:Country [Austria]	1.31	0.98 – 1.74	0.067	1.18	0.90 – 1.55	0.230
Age [65–74 years]:Country [Austria]	1.38	0.92 – 2.06	0.122 ³			
Age [75–84 years]:Country [Austria]	0.97	0.65 – 1.44	0.880 ³			
Age [≥ 85 years]:Country [Austria]	1.02	0.65 – 1.60	0.926 ³			
ICD-10—Mental and Behavioural disorders [yes]:Country [Austria]	0.86	0.63 – 1.18	0.349			
ICD-10—Neoplasms [yes]:Country [Austria]	0.81	0.57 – 1.13	0.208			
ICD-10—Diseases of the blood and blood-forming organs [yes]:Country [Austria]	1.08	0.74 – 1.56	0.700			
ICD-10—Certain infectious and parasitic diseases [yes]:Country [Austria]	0.80	0.51 – 1.25	0.322			
ICD-10—Factors influencing health status [yes]:Country [Austria]	1.03	0.68 – 1.58	0.878			
ICD-10—Diseases of the nervous system [yes]:Country [Austria]	0.98	0.70 – 1.38	0.904			
ICD-10—Endocrine, nutritional and metabolic diseases [yes]:Country [Austria]	0.94	0.72 – 1.23	0.641			
ICD-10—Diseases of the musculoskeletal system [yes]:Country [Austria]	0.93	0.69 – 1.25	0.623			
Surgical procedure within 14 days prior to measurement [yes]:Country [Austria]	0.70	0.48 – 1.03	0.069	0.73	0.50 – 1.07	0.104
Sex [female]:Country [Austria]	0.83	0.63 – 1.09	0.178			
ICD-10—Diseases of the ear [yes]:Country [Austria]	0.92	0.31 – 2.75	0.882			
Random Effects						
σ ² [residual variance]	3.29			3.29		
τ ₀₀ [variability in hospital intercepts]	0.09			0.09		
N [hospitals]	176			176		
Observations	43,984			43,984		
AIC	11,871			11,855		

CDS Care dependency scale, ICD-10 International statistical classification of diseases and related health problems 10th revision, Ref. reference category, significant OR are highlighted in bold

¹ “Model 1” describes the inpatient fall risk adjustment model including all interaction terms regarding patient-related fall risk factors and country and “Model 2” describes the inpatient fall risk adjustment model including the relevant interaction effects found in “Model 1”

² “Model 1” without the interaction term CDS*Country revealed a lower AIC value of 11,869. The interaction term CDS*Country is therefore not included in “Model 2”

³ “Model 1” without the interaction term Age*Country revealed a lower AIC value of 11,870. The interaction term Age*Country is therefore not included in “Model 2”

than 0.16. The effect of sedatives/psychotropic medication intake (OR 1.31, 95% CI 0.98–1.74, $p=0.067$) as well as the effect of having a surgical procedure (OR 0.70, 95% CI 0.48–1.03, $p=0.069$) on the risk of falling in hospital differs between Swiss and Austrian patients. All other interaction terms exceed a p -value of 0.16. Therefore, the odds of experiencing an inpatient fall while, e.g., affected by a mental or behavioural disorder is not dependent on being a patient in Switzerland or Austria (OR 0.86, 95% CI 0.63–1.18, $p=0.349$). In other words, most of the patient-related fall risk variables in the model show a constant risk relationship with the outcome inpatient fall, regardless of the country.

The comparison of “Model 1” and “Model 2” shows that the reduced “Model 2” fits the data better with an AIC value of 11,855 compared with a value of 11,871 for “Model 1” (see Table 3). The three interaction effects remaining in “Model 2” are not statistically significant. If the individual subgroups are examined in detail, both significant and non-significant differences between the countries become apparent. Swiss and Austrian patients without a fall in the last 12 months, without sedative/psychotropic medication or without surgery have no significantly higher or lower odds of falling in hospital (OR 1.19, 95% CI 0.89–1.60, $p=0.250$). However, patients who have fallen in the last 12 months (OR 1.49,¹ 95% CI 1.10–2.01, $p=0.009$) or take sedatives/psychotropic medication (OR 1.40¹, 95% CI 1.05–1.87, $p=0.022$) have significantly higher odds of falling in hospital in Austria than in Switzerland. Although not statistically significant, the risk of falling is lower for Austrian patients with surgery (OR 0.87¹, 95% CI 0.58–1.31, $p=0.508$) than for Swiss patients.

Discriminatory performance of the inpatient fall risk adjustment model

The inpatient fall risk adjustment model was originally developed based on the Swiss data. Applying the model to this data set showed an area under the curve (AUC) of 79.0% (95% CI 77.8–80.2). In comparison (see Fig. 1), the AUC based on the Austrian data set showed a value of 79.9% (95% CI 77.6–82.3) and thus does not differ significantly ($p=0.505$). The covariates included in the inpatient fall risk adjustment model are equally able to predict a fall in hospital in both countries.

After controlling for differences in patient-related fall risk factors between the two countries by applying the inpatient fall risk adjustment model, a statistically significant country difference in the odds of falling in

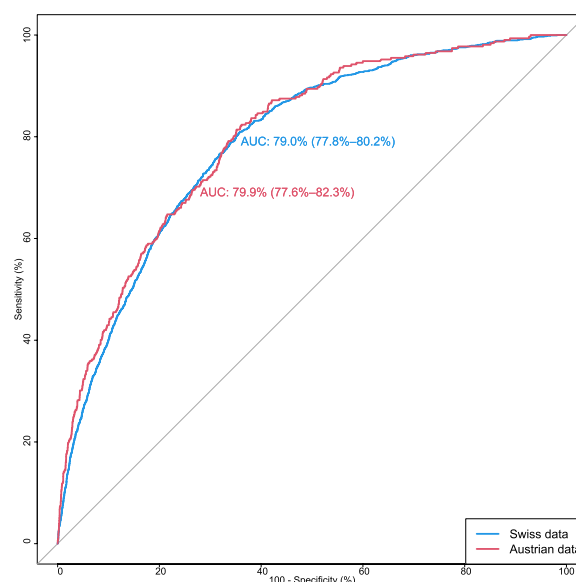


Fig. 1 ROC curves of the inpatient fall risk adjustment model based on Swiss and Austrian data

hospital emerges (Table 4). After risk adjustment, Austrian patients show significantly higher odds of falling in hospital than do Swiss patients (OR 1.38, 95% CI 1.13–1.68, $p=0.002$).

Discussion

In this study, we investigated the extent to which the effects of the patient-related fall risk factors used for risk adjustment in connection with inpatient falls developed with Swiss data remained constant when applied to an external Austrian data set, thus allowing them to be used for a risk-adjusted comparison of inpatient falls between Switzerland and Austria. Of the total of 15 patient-related fall risk variables in the model, we could not find any evidence of relevant interaction effects with the grouping variable “country” for 12 variables. We therefore conclude that the majority of risk factors used in the inpatient fall risk adjustment model are associated with inpatient falls regardless of country, which indicates that the risk factors can safely be used in the risk adjustment model to compare the inpatient falls of Austria and Switzerland [22]. In addition, we demonstrated that the risk of falling in hospital in Switzerland and Austria differs statistically significantly after risk adjustment.

Three of the patient-related fall risk variables included in the models, namely “fall in the last 12 months”, “sedatives/psychotropic medication intake” and “surgical procedure within 14 days prior to measurement”, showed, at least in part, evidence for a non-constant relationship. Even though all three interaction effects in “Model 2” (reduced model, controlling only for relevant country-specific interaction effects identified in “Model 1”) were not statistically significant, they suggest that the relationship between these risk factors and the outcome may differ between the two countries.

¹ The odds ratios of subgroup comparisons, which are not shown in the tables due to the reference category used, were calculated and reported in the text as follows: $OR_{Country}$ multiplied by $OR_{Interaction\ term}$ = $OR_{reported\ in\ text}$.

Table 4 Risk-adjusted comparison of the odds of falling in hospital in Switzerland or Austria

Predictors	OR	95% CI	p-value
CDS [care independent (70–75)]	Ref		
CDS [to a great extent independent (60–69)]	2.26	1.94 – 2.62	< 0.001
CDS [partially dependent (45–59)]	2.98	2.54 – 3.49	< 0.001
CDS [to a great extent dependent (25–44)]	3.32	2.75 – 4.00	< 0.001
CDS [completely dependent (15–24)]	1.97	1.44 – 2.69	< 0.001
Fall in the last 12 months [yes]	2.31	2.06 – 2.58	< 0.001
Sedative/psychotropic medication intake [yes]	1.84	1.64 – 2.07	< 0.001
Age [18–64 years]	Ref		
Age [65–74 years]	1.45	1.23 – 1.71	< 0.001
Age [75–84 years]	1.59	1.36 – 1.86	< 0.001
Age [≥ 85 years]	1.56	1.31 – 1.85	< 0.001
ICD-10—Mental and Behavioural disorders [yes]	1.51	1.34 – 1.70	< 0.001
ICD-10—Neoplasms [yes]	1.21	1.07 – 1.37	0.003
ICD-10—Diseases of the blood and blood-forming organs [yes]	1.25	1.10 – 1.42	0.001
ICD-10—Certain infectious and parasitic diseases [yes]	1.15	0.99 – 1.32	0.060
ICD-10—Factors influencing health status [yes]	1.22	1.03 – 1.45	0.021
ICD-10—Diseases of the nervous system [yes]	1.14	1.00 – 1.30	0.049
ICD-10—Endocrine, nutritional and metabolic diseases [yes]	1.12	1.01 – 1.25	0.040
ICD-10—Diseases of the musculoskeletal system [yes]	0.90	0.80 – 1.00	0.053
Surgical procedure within 14 days prior to measurement [yes]	0.78	0.69 – 0.88	< 0.001
Sex [female]	0.76	0.68 – 0.85	< 0.001
ICD-10—Diseases of the ear [yes]	0.68	0.49 – 0.96	0.029
Country [Austria]	1.38	1.13 – 1.68	0.002
Random Effects			
σ^2 [residual variance]	3.29		
τ_{00} [variability in hospital intercepts]	0.09		
N [hospitals]	176		
Observations	43,984		
AIC	11,856.5		

CDS Care dependency scale, ICD-10 International statistical classification of diseases and related health problems 10th revision, significant OR are highlighted in bold

exceeded the statistical significance level of a p -value smaller than 0.05, significant subgroup differences were found between Switzerland and Austria for two of the risk variables. Specifically, patients in Austria with a fall in the last 12 months or who take sedative/psychotropic medication have a higher risk of falling in hospital than have Swiss patients.

In general, significant interactions and thus non-constant risk relationships according to Moger and Peltola [1], Mohammed et al. [20], Goodacre et al. [21] may be due to the following scenarios: (a) a non-uniform measurement procedure is used (the risk variable is recorded or collected differently in the two countries, which may result in underreporting in one country, for example); (b) the risk variable actually reflects a different risk association with outcome in different patient populations (e.g., the risk of inpatient falls is lower in a group of completely immobile patients who have fallen in the last 12 months

than in a group of mobile patients who have fallen in the last 12 months); c) the quality of care differs between countries in relation to patients with or without risk (e.g., patients with a fall in the last 12 months may or may not receive standard fall prevention measures in one country).

In the context of a fair international comparison, scenarios a and b are problematic because the risk adjustment model assumes a constant risk relationship between the risk variables and the outcome in different populations, which may not actually be the case. This can lead to a risk-adjusted comparison being unintentionally more biased than would be the case with a purely descriptive comparison [20, 22, 49]. Scenario c is unproblematic, as the interaction effect reflects the actual differences in the quality of care provided in the two countries and not a non-constant risk relationship per se. Thus, by assuming a constant risk relationship in the risk adjustment

process, the country which, for example, implements standardised fall prevention measures (quality of care) for high-risk patients (falls in the last 12 months before hospital admission) and thus achieves lower fall rates in this group, is correctly attributed with a higher quality of care.

Because there is no standard procedure or solution for identifying the underlying reasons for an interaction, and because reasons are usually not preclusive [1, 20], we can only speculate about the reasons for the subgroup differences found in our study and thus whether and, if so, to what extent these are problematic for a risk-adjusted comparison of inpatient falls in Switzerland and Austria. Since the measurement in both countries is conducted using the same method, i.e., according to a highly standardised procedure, a non-uniform measurement procedure is rather unlikely to be the reason. Furthermore, the risk factors “fall in the last 12 months” and “sedatives/psychotropic medication intake” are consistently described in the international literature as central risk factors for inpatient falls in various studies done in different contexts as well as in international guidelines [28, 50–55]. It can be assumed that they are relatively consistently associated with a higher risk of falling in hospital, regardless of the country or the population being compared. The subgroup differences found in our study can thus possibly be explained by country-specific differences in fall prevention practice (quality of care) in these subgroups and not by non-constant risk relationships.

In order to be able to substantiate or refute the speculations, a more comprehensive application of the risk adjustment model to international data sets from several countries is recommended. A basic prerequisite is that the hospitals and countries included in the international comparison ensure and adhere to a uniform approach to data collection as well as the precise definition and operationalisation of key variables in order to exclude or at least minimise systematic bias in the underlying data [3, 13, 14, 16].

Interestingly, our study also showed that, based on our sample, the non-risk-adjusted country comparison shows no difference and the comparison adjusted for patient-related fall risk factors shows a significant difference in the odds of falling in hospital between countries. The conclusions that one might draw based on the non-risk-adjusted results are in some ways the opposite of those based on the risk-adjusted results. This highlights the potential pitfalls of a non-risk-adjusted comparison of outcomes across countries, as in a worst-case scenario this could potentially lead to misguided inactivity or actionism by national-level decision-makers and healthcare providers, but also by healthcare workers. Based on the non-risk-adjusted results, it is conceivable

that inpatient fall prevention in Austria would be defined as a non-priority area for quality improvement, therefore shifting the allocation of corresponding resources to other areas, which in turn would not be available for improving fall prevention, especially given the limited resources in the health care system.

It should be noted that the results of an international comparison, whether risk-adjusted or not, should always be contextualised. International comparisons are useful for identifying signals of quality deficits. They do not usually provide an answer to the question of how to improve the quality of care, but offer an evidence-based starting point for a more in-depth analysis to better understand the results found and to stimulate a debate on quality of care and cross-national learning [3, 56]. As a starting point for mutual cross-national learning and to inform clinical practice, it would be interesting for Austrian healthcare institutions to examine what is done differently in practice in Switzerland. In particular, it would be interesting for nurses to analyse which preventive nursing measures are routinely implemented in clinical practice in Switzerland to prevent falls in patients who have had a fall in the last 12 months or who are taking sedative/psychotropic medication.

Strength and limitations

A strength of our study is the first validation and application of an inpatient fall risk adjustment model based on a large sample from two geographically close European countries. In contrast to other international comparisons, which are usually based on administrative data and therefore on data of potentially limited quality, e.g., due to non-uniform coding practices [57], we were able to draw on standardised and uniformly collected primary data. This aspect in particular is highly relevant because, as described in the introduction, differences in an outcome between groups can also be caused by differences in data collection and, consequently, data quality. This possibility could be largely eliminated in our study through the use of uniformly collected data.

It should be noted, however, that in contrast to hospitals in Switzerland, which were obliged to participate in the measurement, and the costs were covered by ANQ, the hospitals in Austria had to pay a fee to participate. A selection bias in Austria needs to be assumed. It is likely that only those hospitals that consider the measurement of quality indicators to be important and are sensitised to it participated in the measurement and invested money accordingly. The differences found in the sample characteristics of the two countries could be an indication of a possible selection bias. In this context, it would have been interesting to compare the ward types participating in the measurement per

country, as an over- or under-representation of certain patient characteristics could be related to an over- or under-representation of certain ward types in the sample. Unfortunately, not all ward types were defined uniformly in Switzerland and Austria, and the individual wards of a hospital sometimes changed their ward type over the measurement years, so that it was not always possible to assign an individual ward to a ward type with certainty. The voluntary nature of participation in the measurement in Austria could also be the reason for the differences in sample size between the two countries. For the risk-adjusted comparison of inpatient fall rates, this is in principle unproblematic, as we were able to show that the effects of the patient-related risk variables are largely constant in the two countries. Accordingly, it is irrelevant how many cases a country contributes to the overall data set [1]. However, due to the before mentioned possible selection bias, the data on which the risk-adjusted comparison of inpatient fall rates is based may not be completely representative for Austria and therefore the generalisability of the results to the national level is limited.

Additionally, in order to increase the stability and the estimates of the risk adjustment model, it would have been desirable to collect data over a longer time period instead of cross-sectional data. This is hardly feasible in terms of resources within the framework of primary data collection at the national level. We tried to take this into account by pooling the three measurement points. Although the LPZ measurement is also carried out in other countries such as the Netherlands and Turkey, it was not possible to include these countries in the analyses due to too few participating hospitals and correspondingly fewer participating patients in these countries. If data from several countries were available in the future, this would favour further external validation of the inpatient fall risk adjustment model.

Conclusions

In our study, we presented first results on the applicability of the inpatient fall risk adjustment model for international comparison. The majority of the patient-related fall risk factors considered in the model can be used for the risk adjustment of inpatient fall rates at the international level, as hardly any significant interactions could be identified. Moreover, the patient-related fall risk factors considered in the inpatient fall risk adjustment model are equally able to predict a fall in hospital, regardless of the country. However, further investigation in a broader international context involving several countries is warranted. In particular, it is

recommended to check to what extent the fall risk factors “fall in the last 12 months” and “sedatives/psychotropic medication intake” interact with the grouping variable country. Furthermore, we were able to show that the results differ when the inpatient fall risk is compared in Switzerland and Austria without and with adjustment for patient-related risk factors. The non-risk-adjusted results imply a non-significant difference in national fall rates between the two countries, but the risk-adjusted comparison suggests a significant one. This illustrates that the application of a risk adjustment model can significantly alter foundations for decision-making and thus may lead to different conclusions by decision-makers regarding quality improvement potential. We conclude that our inpatient fall risk adjustment model enables a risk-adjusted international comparison of inpatient fall rates and thus contributes significantly to achieving the goal of a fair comparison. Also, recognising the imperfection of any risk-adjustment model, we recommend that decision-makers consider risk-adjusted international comparisons whenever possible and available, ideally in addition to national results over time, to incorporate the most comprehensive and robust information possible in assessing potential for improvement in the area of inpatient falls.

Abbreviations

95% CI	95% Confidence interval
AIC	Akaike Information Criterion
AUC	Area Under the Curve
CDS	Care Dependency Scale
ICD-10	International Statistical Classification of Diseases and Related Health Problems 10th Revision
LPZ	National Prevalence Measurement of Quality of Care (in Dutch: Landelijke Prevalentiemeting Zorgkwaliteit)
OR	Odds Ratio
ROC	Receiver Operating Characteristics

Supplementary Information

The online version contains supplementary material available at <https://doi.org/10.1186/s12913-024-10839-x>.

Supplementary Material 1.

Supplementary Material 2.

Supplementary Material 3.

Acknowledgements

We would like to thank Prof. Dr. Marianne Müller for her advice on statistical issues and her support during the writing process.

Authors' contributions

Conceptualization of the measurement, NB, IE, SH, SB, JS; conceptualization of the manuscript, NB; data collection and curation, NB, SB; data analysis, NB; data interpretation, NB, IE, SH, SB, JS; writing – original draft preparation, NB; writing – review and editing, IE, SH, SB, JS. All authors read and approved the final manuscript.

Funding

The annual quality measurement in Switzerland's acute care hospitals and thus the data collection for this study was financed by the Swiss National Association for Quality Development in Hospitals and Clinics (ANQ). In Austria, the data collection was financed by the participating hospitals. The processing and analysis of the data used in this study was funded by Bern University of Applied Sciences.

Availability of data and materials

The Swiss and Austrian data that support the findings of this study are available from the Swiss National Association for Quality Development in Hospitals and Clinics (ANQ) and the Department of Nursing Science at the Medical University of Graz, respectively, but restrictions apply to the availability of these data, which were used under license for the current study, and so are not publicly available. Therefore, the data are available upon reasonable request and with the permission of the Swiss Association for Quality Development in Hospitals and Clinics (ANQ) and the Department of Nursing Science at the Medical University of Graz from the corresponding author.

Declarations

Ethics approval and consent to participate

This study was conducted in accordance with the ethical principles of the Declaration of Helsinki and with the approval of the following ethics committees: In Austria the study was approved by the medical ethics committee of the Medical University of Graz (Ek-Nr 20–192 ex 08/09); in Switzerland Swiss ethics and the cantonal ethics committees had classified the measurement in 2012 as a quality measurement and therefore declared that the LPZ measurement is not within their purview, as it does not fall under the Swiss Human Research Act (see Bernet et al. [28] for further details). In addition, the Ethics Advisory Board of the Bern University of Applied Sciences approved in 2022 that the anonymized data from the LPZ measurement may be further used without further approval at a cantonal ethics committee (Ref. EAB2022_012). All patients or their legal representatives received written information about the measurement. Written (in Austria) and verbal (in Switzerland) informed consent was obtained from all participants or their legal representatives. Verbal informed consent was authorised in Switzerland by Swiss ethics (Swiss Association of Research Ethics Committees) and the cantonal ethics committees, as described in detail in Bernet et al. [28]. Patients were informed that they could withdraw their consent at any time without justification and without consequences.

Consent for publication

Not applicable.

Competing interests

The authors declare no competing interests.

Author details

¹School of Health Professions, Applied Research & Development in Nursing, Bern University of Applied Sciences, Murtenstrasse 10, Bern 3008, Switzerland.

²Department of Health Services Research, Care and Public Health Research Institute, Maastricht University, PO BOX 616, Maastricht 6200 MD, the Netherlands. ³Institute of Nursing Science, Medical University of Graz, Neue Stiftingtalstraße 6/P06-WEST, 8010 Graz, Austria.

Received: 8 May 2023 Accepted: 7 March 2024

Published online: 13 March 2024

References

- Moger TA, Peltola M. Risk adjustment of health-care performance measures in a multinational register-based study: A pragmatic approach to a complicated topic. *SAGE Open Medicine*. 2014;2:2050312114526589. <https://doi.org/10.1177/2050312114526589>.
- Kiivet R, Sund R, Linna M, Silverman B, Pisarev H, Friedman N. Methodological challenges in international performance measurement using patient-level administrative data. *Health Policy*. 2013;112(1):110–21. <https://doi.org/10.1016/j.healthpol.2013.05.011>.
- Smith PC, Mossialos E, Papanicolas I, Leatherman S. Performance Measurement for Health System Improvement: Experiences, Challenges and Prospects. Cambridge: Cambridge University Press; 2010.
- OECD. Improving Value in Health Care. Paris: OECD Publishing; 2010. Available from: <https://doi.org/10.1787/9789264094819-en>. Accessed 12 Mar 2024.
- Smith PC. Getting the most out of international comparison. 2015. <https://www.nuffieldtrust.org.uk/news-item/getting-the-most-out-of-international-comparison>. Accessed on 02 Feb 2023.
- Carinci F, Van Gool K, Mainz J, Veillard J, Pichora EC, Januel JM, et al. Towards actionable international comparisons of health system performance: expert revision of the OECD framework and quality indicators. *Int J Qual Health Care*. 2015;27(2):137–46. <https://doi.org/10.1093/intqhc/mzw004>.
- OECD. Health Care Quality Framework. n.d. <https://www.oecd.org/els/health-systems/health-care-quality-framework.htm>. Accessed on 31 Jan 2024.
- Rotar AM, van den Berg MJ, Kringos DS, Klazinga NS. Reporting and use of the OECD Health Care Quality Indicators at national and regional level in 15 countries. *Int J Qual Health Care*. 2016;28(3):398–404. <https://doi.org/10.1093/intqhc/mzw027>.
- OECD. OECDStat. <https://stats.oecd.org/Index.aspx?ThemeTreeId=9>. Accessed on 01 Feb 2024.
- OECD. Health at a Glance 2023. Paris: OECD Publishing; 2023. Available from: <https://www.oecd-ilibrary.org/content/publication/7a7afb35-en>.
- Busse R, Klazinga N, Panteli D, Quentin W. Improving healthcare quality in Europe: Characteristics, effectiveness and implementation of different strategies. 2019. <https://apps.who.int/iris/bitstream/handle/10665/327356/9789289051750-eng.pdf?sequence=1&isAllowed=y>. Accessed on 17 May 2021.
- Williamson PR, Barrington H, Blazeby JM, Clarke M, Gargon E, Gorst S, et al. Review finds core outcome set uptake in new studies and systematic reviews needs. *J Clin Epidemiol*. 2022;150:154–64. <https://doi.org/10.1016/j.jclinepi.2022.06.016>.
- Wait S, Nolte E. Benchmarking health systems trends conceptual issues and future perspectives. *Benchmarking*. 2005;12(5):436–48. <https://doi.org/10.1108/14635770510619366>.
- Bernet NS, Thomann S, Kurpicz-Briki M, Roos L, Everink IHJ, Schols J, et al. Potential of Electronic Medical Record Data for National Quality Measurement. *Studies in Health Technology and Informatics*. 2022;292:51–6. <https://doi.org/10.3233/shti220320>.
- Lilford R, Mohammed MA, Spiegelhalter D, Thomson R. Use and misuse of process and outcome data in managing performance of acute medical care: avoiding institutional stigma. *The Lancet*. 2004;363(9415):1147–54. [https://doi.org/10.1016/S0140-6736\(04\)15901-1](https://doi.org/10.1016/S0140-6736(04)15901-1).
- Forde I, Morgan D, Klazinga NS. Resolving the challenges in the international comparison of health systems: The must do's and the trade-offs. *Health Policy*. 2013;112(1):4–8. <https://doi.org/10.1016/j.healthpol.2013.01.018>.
- Oderkirk J, Ronchi E, Klazinga N. International comparisons of health system performance among OECD countries: Opportunities and data privacy protection challenges. *Health Policy*. 2013;112(1):9–18. <https://doi.org/10.1016/j.healthpol.2013.06.006>.
- Lane-Fall MB, Neuman MD. Outcomes measures and risk adjustment. *Int Anesthesiol Clin*. 2013;51(4):10–21. <https://doi.org/10.1097/AIA.0b013e3182a70a52>.
- Ash A, Fienberg S, Louis T, Normand S-L, Stukel T, Utts J. Statistical Issues in Assessing Hospital Performance. 2012. <https://www.cms.gov/medicare/quality-initiatives-patient-assessment-instruments/hospitalqualityinits/downloads/statistical-issues-in-assessing-hospital-performance.pdf>. Accessed on 02 Feb 2023.
- Mohammed MA, Deeks JJ, Gillingham A, Rudge G, Carmalt M, Stevens AJ, et al. Evidence of methodological bias in hospital standardised mortality ratios: retrospective database study of English hospitals. *BMJ*. 2009;338: b780. <https://doi.org/10.1136/bmj.b780>.
- Goodacre S, Wilson R, Shephard N, Nicholl J. Derivation and validation of a risk adjustment model for predicting seven day mortality in emergency medical admissions: mixed prospective and retrospective cohort study. *BMJ*. 2012;344: e2904. <https://doi.org/10.1136/bmj.e2904>.
- Nicholl J. Case-mix adjustment in non-randomised observational evaluations: the constant risk fallacy. *J Epidemiol Community Health*. 2007;61(11):1010–3. <https://doi.org/10.1136/jech.2007.061747>.

23. Dubois C-A, D'amour D, Brault I, Dallaire C, Déry J, Duhoux A, et al. Which priority indicators to use to evaluate nursing care performance? A discussion paper *J Adv Nurs*. 2017;73(12):3154–67. <https://doi.org/10.1111/jan.13373>.
24. Koch D, Kutz A, Conca A, Wenke J, Schuetz P, Mueller B. The relevance, feasibility and benchmarking of nursing quality indicators: A Delphi study. *J Adv Nurs*. 2020;76(12):3483–94. <https://doi.org/10.1111/jan.14560>.
25. Slawomirski L, Klazinga N. The economics of patient safety: From analysis to action. 2020. <https://www.oecd.org/health/health-systems/Economics-of-Patient-Safety-October-2020.pdf>. Accessed on 13 Oct 2021.
26. Commission J. Preventing falls and fall-related injuries in health care facilities. Sentinel Event Alert. 2015;55:1–5.
27. Morello RT, Barker AL, Watts JJ, Haines T, Zavarsek SS, Hill KD, et al. The extra resource burden of in-hospital falls: a cost of falls study. *Med J Aust*. 2015;203(9):367. <https://doi.org/10.5694/mja15.00296>.
28. Bernet NS, Everink IHJ, Schols JMGA, Halfens RJG, Richter D, Hahn S. Hospital performance comparison of inpatient fall rates; the impact of risk adjusting for patient-related factors: a multicentre cross-sectional survey. *BMC Health Serv Res*. 2022;22(1):225. <https://doi.org/10.1186/s12913-022-07638-7>.
29. Van Nie NC, Schols JMGA, Meesterberends E, Lohrmann C, Meijers JMM, Halfens RJG. An international prevalence measurement of care problems: study protocol. *J Adv Nurs*. 2013;69(9):c18–29. <https://doi.org/10.1111/jan.12190>.
30. Benchimol EI, Smeeth L, Guttman A, Harron K, Hemkens LG, Moher D, et al. Das RECORD-Statement zum Berichten von Beobachtungsstudien, die routinemäßig gesammelte Gesundheitsdaten verwenden. *Z Evid Fortbild Qual Gesundhwes*. 2016;115–116:33–48. <https://doi.org/10.1016/j.zefq.2016.07.010>.
31. Benchimol EI, Smeeth L, Guttman A, Harron K, Moher D, Petersen I, et al. The REporting of studies Conducted using Observational Routinely-collected health Data (RECORD) Statement. *PLOS Med*. 2015;12(10): e1001885. <https://doi.org/10.1371/journal.pmed.1001885>.
32. Thomann S, Rösli R, Richter D, Bernet NS. Auswertungskonzept ANQ: Nationale Prävalenzmessung Sturz und Dekubitus. 2020. https://www.anq.ch/wp-content/uploads/2017/12/ANQ_Sturz_Dekubitus_Auswertungskonzept.pdf. Accessed on 14 Dec 2021.
33. Kellogg International Work Group on the Prevention of Falls by the Elderly. The prevention of falls in later life. A report of the Kellogg International Work Group on the Prevention of Falls by the Elderly. *Dan Med Bull*. 1987;34(Supplement 4):1–24.
34. Dijkstra A, Smith J, White M. Manual Care Dependency Scale. Measuring care dependency with the Care Dependency Scale (CDS). A manual. 2006. <https://www.care2share.eu/dbfiles/download/29>. Accessed on 02 Dec 2019.
35. Lohrmann C, Dijkstra A, Dassen T. The Care Dependency Scale: an assessment instrument for elderly patients in German hospitals. *Geriatr Nurs*. 2003. <https://doi.org/10.1067/mgn.2003.8>.
36. Eglseer D, Halfens RJG, Schols JMGA, Lohrmann C. Dysphagia in Hospitalized Older Patients: Associated Factors and Nutritional Interventions. *J Nutr, Health and Aging*. 2018;22(1):103–10. <https://doi.org/10.1007/s12603-017-0928-x>.
37. Schüttengruber G, Halfens RJG, Lohrmann C. Care dependency of patients and residents at the end of life: A secondary data analysis of data from a cross-sectional study in hospitals and geriatric institutions. *J Clin Nurs*. 2022;31(5–6):657–68. <https://doi.org/10.1111/jocn.15925>.
38. World Health Organization. International Statistical Classification of Diseases and Related Health Problems 10th Revision. 2016. <https://icd.who.int/browse10/2016/en>. Accessed on 19 Jan 2023.
39. Clark M. Michael Clark: Convergence Problems. 2020. <https://m-clark.github.io/posts/2020-03-16-convergence/>. Accessed on 20 Dec 2022.
40. Heikkilä A, Lehtonen L, Junttila K. Fall rates by specialties and risk factors for falls in acute hospital: A retrospective study. *J Clin Nurs*. 2023;32(15–16):4868–77. <https://doi.org/10.1111/jocn.16594>.
41. Thomann S, Schlunegger M, Rösli R, Kammer L, Riedweg J, Richter D, et al. Sturz und Dekubitus Akutsomatik Erwachsene - Nationaler Vergleichsbericht Messung 2019. 2020. https://www.anq.ch/wp-content/uploads/2020/09/ANQakut_Sturz_Dekubitus_Erwachsene_Nationaler-Vergleichsbericht_2019.pdf. Accessed on 01 Feb 2024.
42. Heinze G, Dunkler D. Five myths about variable selection. *Transpl Int*. 2017;30(1):6–10. <https://doi.org/10.1111/tri.12895>.
43. R Core Team. R: A Language and Environment for Statistical Computing. 2021. Available from: <https://www.R-project.org/>. Accessed 12 Mar 2024.
44. Bates D, Mächler M, Bolker B, Walker S. Fitting Linear Mixed-Effects Models Using lme4. *J Stat Softw*. 2015;67(1):1–48. <https://doi.org/10.18637/jss.v067.i01>.
45. Lüdtke D. sjPlot: Data Visualization for Statistics in Social Science. 2021. Available from: <https://CRAN.R-project.org/package=sjPlot>. Accessed 12 Mar 2024.
46. Fox J. Effect Displays in R for Generalised Linear Models. *J Stat Softw*. 2003;8(15):1–27. <https://doi.org/10.18637/jss.v008.i15>.
47. Fox J, Weisberg S. An R Companion to Applied Regression. 3rd ed. Thousand Oaks CA: Sage; 2019.
48. Robin X, Turck N, Hainard A, Tiberti N, Lisacek F, Sanchez J-C, et al. pROC: an open-source package for R and S+ to analyze and compare ROC curves. *BMC Bioinformatics*. 2011;12(1):77. <https://doi.org/10.1186/1471-2105-12-77>.
49. Deeks JJ, Dinnes J, D'Amico R, Sakaravitch C, Song F, et al. Evaluating non-randomised intervention studies. *Health Technol Assess*. 2003;7(27):iii–x, 1–173. <https://doi.org/10.3310/hta7270>.
50. Clay F, Yap G, Melder A. Risk factors for in hospital falls: Evidence Review. 2018. https://monashhealth.org/wp-content/uploads/2019/01/Risk-factors-for-falls_Final-27082018.pdf. Accessed on 06 June 2021.
51. Deandrea S, Bravi F, Turati F, Lucenteforte E, La Vecchia C, Negri E. Risk factors for falls in older people in nursing homes and hospitals. A systematic review and meta-analysis. *Arch Gerontol Geriatr*. 2013;56(3):407–15. <https://doi.org/10.1016/j.archger.2012.12.006>.
52. Hoedl M, Eglseer D, Bernet NS, Everink I, Gordon AL, Lohrmann C, et al. Which factors influence the prevalence of institution-acquired falls? Results from an international, multi-center, cross-sectional survey. *J Nurs Sch*. 2021;00:1–8. <https://doi.org/10.1111/jnu.12758>.
53. Ishikuro M, Ramón Gutiérrez Ubeda S, Obara T, Saga T, Tanaka N, Oikawa C, et al. Exploring risk factors of patient falls: a retrospective hospital record study in Japan. *Tohoku J Exp Med*. 2017;243(3):195–203. <https://doi.org/10.1620/tjem.243.195>.
54. National Institute for Health and Care Excellence [NICE]. Assessment and prevention of falls in older people. 2013. <https://www.nice.org.uk/guidance/CG161>. Accessed on 15 Apr 2021.
55. Registered Nurses' Association of Ontario. Preventing Falls and Reducing Injury from Falls. 2017. https://rnao.ca/sites/rnao-ca/files/bpg/FALL_PREVENTION_WEB_1207-17.pdf. Accessed on 07 June 2021.
56. Kossarova L, Blunt I, Bardsley M. Focus on: International comparisons of healthcare quality. 2015. https://www.nuffieldtrust.org.uk/files/2018-10/1540242820_qualitywatch-international-comparisons-disclaimer-update.pdf. Accessed on 02 Feb 2023.
57. Papanicolas I, Smith P. Health System Performance Comparison: An Agenda for Policy, Information and Research. 1st ed. Maidenhead: Open University Press; 2013.

Publisher's Note

Springer Nature remains neutral with regard to jurisdictional claims in published maps and institutional affiliations.